Title: The limits of direct community modeling approaches for broad-scale
predictions of ecological assemblage structure

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Keywords: community assembly, stacked-species distribution models,
macroecological community model, dietary traits, species prevalence,
biodiversity.

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Type of paper: Research paper

The number of words in main body: 3,849

25 **ABSTRACT**

26 Two distinct modeling approaches are often used when predicting biodiversity
27 patterns: stacking of species models (predict first, assemble later), and directly
28 modeling a characteristic of a community such as species richness (assemble first,
29 predict later). The relative utility of these two approaches for biogeographic,
30 macroecological and global change analyses is uncertain. Here we compared the two
31 approaches by predicting current-day avian dietary guild structure of assemblages
32 worldwide. We found that the stacked-species modeling approach consistently
33 predicted the geographic distribution of observed dietary guilds better than a direct
34 community modeling approach. The exception was for plant-eating birds, especially
35 frugivores, which are expected to have particularly strong climatic constraints on their
36 diversity and distributions. Assemblage-level biodiversity patterns predicted by
37 community-based modeling approaches, such as the stacked-species and direct
38 community modeling approaches in this study, offer a means to help guide
39 conservation decisions for determining environmental suitability and analyzing
40 diversity hotspots. However, our results generally caution against the widespread use
41 of direct community modeling approach at the large spatial extents for predicting
42 species assemblages.

1. Introduction

A major focus in biogeography and macroecology is to accurately model and predict geographic variation in the composition and attributes of biological communities. This pursuit is particularly relevant for projecting how communities and their many functions may be perturbed under impending climate change (Jetz et al., 2007, Dawson et al., 2011, Bellard et al., 2012). The potential constraints shown on community structure arising from assembly rules limiting species' coexistence have long been realized (Weiher and Keddy, 1999, Ackerly and Cornwell, 2007). But, how such constraints impinge on our ability to predict community structure and ecosystem functioning and how consistently they are shown at different spatial scales, especially in a broad-scale climate change future, still remain uncertain (Thuiller et al., 2003, Baselga and Araújo, 2009, Belmaker and Jetz, 2013). If prevalent, assembly rules may under human-driven global change significantly affect the re-aggregation of communities and thus change the individual responses of species. Such community-level constraints on future species distributions would have significant consequences for the appropriate modeling approaches (Araujo and Luoto, 2007, Heikkinen et al., 2007, Baselga and Araújo, 2009).

Communities and their attributes (such as total species richness, functional indices or prevalence of a given group) are commonly modeled using a *stacked-species* community modeling approach that first models the spatial distributions of each species and then derives community attributes at different locations by aggregating the predicted individual species distributions (Guisan and Zimmermann, 2000, Ferrier and Guisan, 2006, Dubuis et al., 2011). This kind of approach offers detailed information on the exact identity of the species in the predicted assemblages and has often led to accurate predictions of community attributes for small and spatially

68 restricted datasets (Leathwick et al., 2006, Elith and Leathwick, 2007, Algar et al.,
69 2009, Baselga and Araújo 2009, Chapman and Purse, 2011, Dubuis et al., 2011).
70 However, by assuming that species assemble independently of each other, this
71 approach ignores potential interrelationships among species comprising a community.
72 One simple alternative is a *direct community* modeling approach that statistically
73 associates aggregate community attributes with environmental variables (Ferrier and
74 Guisan, 2006, Baselga and Araújo, 2009, Dubuis et al., 2011). Following in the
75 footsteps of macroecological models of species richness (Brown, 1981, Jetz and
76 Rahbek, 2002), this approach is straightforward to use with large, multi-species
77 datasets and putatively able to accommodate species interdependencies within
78 communities because it accommodates spatial co-occurrences of species to infer
79 community-level patterns (Baselga and Araújo, 2009, Chapman and Purse, 2011).
80 These factors combined with the approach's ability to quickly synthesize complex
81 data for conservation applications have led to notable popularity (Ferrier and Guisan,
82 2006, Algar et al., 2009, Chapman and Purse, 2011).

83 An increasing number of studies are considering community attributes such as total
84 species richness, proportional richness or community functional attributes as a
85 response variable for current-day predictions and projections under climate change
86 (Lemoine and Bohning-Gaese, 2003, Kissling et al., 2007, Steinmann et al., 2009,
87 Kissling et al., 2010, Sommer et al., 2010, Chown et al., 2012, Dubuis et al., 2013,
88 Venevskaya et al., 2013, D'Amen et al., 2015, Distler et al., 2015). One specific
89 suggestion is that geographic patterns of energetic “capacity constraints”, the limits on
90 species richness imposed by available energy, may be re-arranged under climate
91 change and affect future community assemblages (Sommer et al., 2010, Chown et al.,
92 2012). Biotic associations between taxonomic groups show sensitivity when

forecasting future biodiversity (Kissling et al., 2007, Kissling et al., 2010). Accounting for trait data through an emerging trait biogeographic perspective reveals heterogeneous taxonomic partitioning of assemblages (Kissling et al., 2012, Ko et al., 2014). So far, community-based models are increasingly used to examine attributes of communities.

Despite this popularity, the appeal of potentially addressing biotic constraints in projecting communities under climate change, and an already existing debate about the choice of the *stacked-species* versus *direct community* modeling approach (Ferrier and Guisan, 2006, Leathwick et al., 2006, Baselga and Araújo, 2009, Chapman and Purse, 2011, Dubuis et al., 2011, Guisan and Rahbek, 2011, Distler et al., 2015), empirical evaluations are rare. Algar et al. (2009) found good performance of *direct community* compared to *stacked-species* modeling approaches for predicting changes in butterfly diversity in Canada. And for a study along a Swiss elevational gradient, Dubuis et al. (2011) showed that, while having slightly less predictive ability, only *direct community* modeling approaches were able to recover the specific shape of the elevation-richness relationship (Bonthoux et al., 2013). More recently, Distler et al. (2015) found that *direct community* modeling approaches could provide more accurate estimates of North American bird species richness than *stacked-species* modeling approaches, particularly during the summer season. However, apart from these case studies, the general suitability and performance of the community-level modeling approaches for broader spatial scales and for ecologically partitioned groups remain unclear.

The goal of our study is to provide such a general assessment and resolve the relative appropriateness of different methodologies for modeling patterns of functional community attributes over a broad geographic extent (and the typically

concomitant coarse spatial grain) and a large range of species and communities. We use a global dataset of bird dietary preferences and geographic distributions to compare model performance of the *stacked-species* modeling approach with two types of *direct community* modeling approaches, one with total and the other with relative (or proportional) richness of different dietary species groups as response variable. Relative guild species richness, or guild prevalence, is a community attribute with important functional implications and has previously been shown to exhibit strong environmental associations (Kissling et al., 2012). We here ask whether the two types of *direct community* modeling approaches are indeed able to predict the variation in guild assemblage along climatic gradients and, by extension, are suited to address their potential perturbation under climate change. If weak environmental associations of the response variable or overall limited predictive performance constrain the *direct community* modeling approach, is the *stacked-species* modeling approach able to successfully capture the aggregate response and predict its spatial variation? We here assess this for a range of guild groupings that vary in the respective environmental associations. Given the direct consequences of such changes for the functioning and services of ecosystems (Schmitz et al., 2003, Mooney et al., 2009, Kardol et al., 2010), identifying suitable approaches for modeling these and other community attributes is of strong applied importance and useful for biodiversity conservation and management.

2. Methods

2.1. Species data and environmental predictors

We used a comprehensive database of recent expert-based breeding distributions of

all 9,993 bird species in the world, spanning a latitudinal range of 60°S to 85°N (see Jetz et al., 2012) and mol.org for sources, taxonomic treatment and individual maps). Marine and pelagic species as well as species with less than four occurrence grid cells were excluded from the analysis, leaving a total of 8,472 bird species. We intersected these distributions with a global 110 × 110 km equal area grid (approximately 1° near equator), a spatial resolution that according to recent validation of these sorts of expert-based maps offers sufficiently low false presence rates (Hurlbert and Jetz, 2007). This resulted in 11,079 grid cell assemblages for analysis.

We obtained estimates of the proportional use of each of seven food categories (fruits, nectar, plants, seeds, invertebrates, vertebrates, and carcasses) for each bird species from Wilman et al. (2014) to assign each species to a dietary guild. In a “coarse” dietary classification all species were classified as either primary consumers or high-level consumers based on their main proportional use of summed plant (i.e. fruits, nectar, plants, and seeds) and summed animal (i.e. invertebrates, vertebrates, and scavengers) diets. “Fine” dietary guilds were identified by highest proportional use as frugivores, nectarivores, herbivores, granivores, insectivores, carnivores, or scavengers. Those species exhibiting equal use of multiple summed or individual diets were classified as omnivores (named mixed consumers in the coarse dietary guilds and omnivores in the fine dietary guilds). For similar treatment and additional details see also Ko et al. (2014).

We extracted nine environmental predictor variables (all log₁₀-transformed) including one topographic and eight climatic predictors over the same 110 km grid. All selected variables are known to exert a strong influence on the distributions of individual species as well as overall and guild species richness (Jetz and Rahbek, 2002, Field et al., 2009, Kissling et al., 2012, Ko et al., 2014). Elevational range was

derived from GTOPO30 at 30-arc seconds (approximately 1 km), produced by the USGS (<http://eors.usgs.gov/>). Additionally, mean annual temperature, temperature seasonality (standard deviation of monthly means), mean temperatures of the coldest and warmest months, total annual precipitation, seasonality of precipitation (coefficient of variation of monthly precipitation), and total precipitation in the driest and wettest months were obtained for the 1975-2001 period (representing current conditions, i.e. 2000, in this study) from CRU TS 2.1 ([Mitchell and Jones, 2005](#)) with an original spatial resolution of 30-arc minutes (approximately 50 km). All environmental predictor variables at the coarse scale of 110 km grid cells (i.e. the spatial resolution of species distributions) were expressed by the area-weighted average values of fine scale data (i.e. topographic and climatic predictors in approximately 1 km and 50 km resolutions, respectively).

2.2. Models

Considering the success of ensembles of models in reducing both false negative and positive errors in predictions of species distributions ([Araujo and New, 2007](#)), we used information from ensemble modeling instead of picking a single ‘best’ model in this study. To model species distributions using the *stacked-species* modeling approach, we applied seven different species distribution modeling algorithms ([Elith et al., 2006](#)) to each of the 8,472 bird species, including one classification method (mixture discriminant analysis, MDA), three regression methods (generalized linear models, GLM, generalized additive models, GAM and multivariate adaptive regression splines, maximum entropy, MARS), and three machine learning methods (general boosting method, GBM, random forests, RF and maximum entropy, MAXENT). Among these we selected the four models with the highest AUC (area

under the curve) values for further use (see below for AUC calculation). We weighted the predictions of all these four models according to their pre-evaluated AUC model performance (Marmion et al., 2009) to generate a final probabilistic distribution map for each species. Species presences based on maps of current species distributions and absences within 4,000 km of a species' current distribution represented suitable and unsuitable species locations, respectively. Because absences far from species occurrences are less informative than those closer (VanDerWal et al., 2009), absence grid cells were assigned weights that exponentially decreased with distance from the nearest species presence grid cell and used in all seven modeling algorithms. As sampling prevalence has been shown to influence the accuracy of species distribution models (McPherson et al., 2004), equal numbers of species presences and absences were assigned for model calibration except for species with <10 presence grid cells, which were assigned 100 absences. We first ran all seven algorithms with the nine environmental predictor variables. The variables with the high contributions averaged across those seven algorithms were then left in the final modeling. For species with ≥ 10 presence cells, the presence cells were divided into four quadrants based on their coordinates above or below the median latitude and median longitude of all presence localities (Peterson et al., 2006). A random subset of three of quadrants of species presence and absence data was used to calibrate the model in the first run while the remaining one for evaluation (Peterson and Shaw, 2003, Peterson et al., 2006), using a threshold independent receiver operating characteristic (ROC) curve along with the associated AUC; 100% of the available species presence and (selected) absence data were used for making final predictions. For species with <10 presence cells, we used a jackknife validation to select the best-performing modeling algorithms (Pearson et al., 2006). Seven modeling algorithms were evaluated using the jackknife procedure

and final predictions were obtained by averaging the modeling algorithms which had a corresponding p value less than 0.01 in the jackknife procedure. Predicted guilds and total richness were then given as the sum of predicted presence probabilities for each species in each cell (Calabrese et al., 2014).

For the *direct community* modeling approach, we implemented four modeling algorithms: generalized linear models (GLM), generalized additive models (GAM); general boosting method (GBM) and random forests (RF). GLM, GAM, and GBM were implemented as Poisson regressions for observed guild richness as response and as Bernoulli/binomial regression for guild prevalence as response. For GBM and RF we used a maximum of 2,000 trees. For each modeling algorithm, beginning with all nine environmental predictors, one predictor variable at a time was sequentially removed using a supervised backward stepwise variable selection procedure. The calibration and evaluation datasets were also obtained by randomly splitting the original dataset into four quadrants, using the same partition way as the *stacked-species* modeling approach. The model with the highest adjusted R^2 and minimum number of predictor variables was selected as the best model for each modeling algorithm. The final *direct community* predictions were obtained by weighting the predictions of the four modeling algorithms according to the R^2 of their best models.

2.3. Statistics

We compared the results provided by the *stacked-species* and *direct community* modeling approaches using predicted prevalence (i.e. relative or proportional richness) of three coarse and eight fine dietary guilds. For this we truncated the richness predictions for each grid cell provided by the *stacked-species* and *community richness*

approaches (based on summed probabilities and predicted richness, respectively) to the closest integer values and divided predicted guild richness by predicted total richness. For all modeling approaches, we limited the geographic scope of predictions to cells with non-zero values for the response variable, i.e. those currently occupied by at least one guild member.

We evaluated the relationship between observed and predicted guild prevalence and calculated the t statistic for the difference from a slope equal to 1 (Equation: $t = (S_{op} - 1) / SE$, where S_{op} is the slope of the observed versus predicted regression lines and SE is the standard error of the slope), root mean squared deviation and R^2 value of the observed versus predicted regression lines (Piñeiro et al., 2008). Modeling algorithms were implemented within the “BIOMOD” (Thuiller et al., 2009) and “dismo” R packages and statistical analyses were performed using R version 2.12.0 (R Development Core Team, 2011).

3. Results

Among the 8,472 species analyzed, the seven individual-species level modeling algorithms, RF, GBM, MAXENT and GAM generally showed the highest predictive accuracy for species with ≥ 10 presences, while GLM, MARS, MDA and GBM performed well for species with < 10 presences (Fig. A.1 in Appendix A). Both at the coarse and fine ecological resolutions, the *stacked-species* and *direct community* modeling approaches varied substantially in their ability to predict the observed prevalence, i.e. the relative or proportional richness, of avian dietary guilds worldwide (Table 1, Figs. 1 and 2). The *stacked-species* modeling approach consistently yielded predictions for guild prevalence that were very close to those observed, as indicated

by slope estimates very close to 1 (low absolute t values for a deviation from 1) and a low root mean squared deviations (Table 1). The two types of *direct community* modeling approaches using richness and prevalence as response variable showed mostly very weak performance that was only in some cases close to any of the *stacked-species* results (e.g. frugivores; Table 1). Both *community richness* and *prevalence* modeling approaches were mostly unable to predict the strong spatial variation in guild prevalence and tended to strongly over-predict relative guild richness in low-prevalence assemblages, and under-predict it in high-prevalence assemblages (Figs. 1 and 2).

In the *stacked-species* modeling approach predictive performance was similarly high for all guilds (Fig. 3, Fig. A.2 in Appendix A), with only a slight lower goodness of fit for scavengers and the three coarse guilds. This contrasts strongly with the *direct community* modeling approach, in which fit was strongest for frugivores and herbivores, but then deteriorated rapidly for all other guilds and especially the coarse guild groupings. Fits differed little among the two *direct community* modeling approaches. In no case was their fit superior to that of the *stacked-species* approach, but the *direct community* modeling approach did offer close to comparable strengths for predicting plant-eating birds.

Although greatly variable among the lower-richness assemblages, (“Observation” panels in Figs. 4 and 5), globally observed guild prevalence was surprisingly uniform along the global richness gradient. For example, the global prevalence of primary and higher-level consumers centered around 0.3 and 0.7, respectively. The *stacked-species* modeling approach was able to capture both the variation and magnitude of prevalence along the richness gradient reasonably well, whereas the *direct community* modeling approach misrepresented the observed prevalence, particularly in the low

richness assemblages. These patterns were similar for most fine dietary guilds (Fig. 5), with the exception of plant-eating guilds, especially frugivores, where the guild prevalence strongly increased with increasing total richness. The exceptional pattern for frugivores was also evident in the direct assessment of predicted vs. observed prevalence (Table 1, Figs. 2 and 3), where the two types of *direct community* modeling approaches, i.e. *community richness* and *prevalence* models, also offered a surprisingly reliable prediction.

4. Discussion

Statistical approaches to predict species assemblages in communities are numerous and highly varied. Comparisons between two prominent approaches—*stacked-species* versus *direct community* modeling—have produced mixed conclusions about their relative predictive ability (Ferrier and Guisan, 2006, Leathwick et al., 2006, Baselga and Araújo, 2009, Dubuis et al., 2011, Guisan and Rahbek, 2011, Distler et al., 2015). Our comparison found that the *direct community* modeling approach showed as good performance as the *stacked-species* modeling approach only for frugivores. Previous analyses have concluded that *direct community* modeling approaches would be favored when species share similar environmental associations and face strong assembly constraints (Baselga and Araújo, 2009, Chapman and Purse, 2011, Dubuis et al., 2011). Our analysis does not support this conclusion, despite the fact that the ecological group, i.e. dietary guilds, analyzed in this study share similar environmental conditions and assembly constraints. Instead, the *stacked-species* modeling approach predicted the prevalence of guilds observed today equally or better than the two types of *direct community* modeling approaches. It has been

318 claimed that a drawback of *stacked-species* modeling approach is that it overpredicts
319 species richness, as in the case of analyses of Swiss plants and New World birds at
320 resolutions of 25m and 1° latitude/longitude, respectively ([Guisan and Rahbek, 2011](#)),
321 because the thresholding approach used to relate predicted probabilities to occurrence
322 may lead to bias ([Guillera-Arroita et al., 2015](#)).

323 We offer some reasons why the two types of *direct community* modeling approach
324 may not be as suitable as the *stacked-species* modeling approach. First, the *direct*
325 *community* modeling approach only estimates one aggregate response – environment
326 association for all species in a grouping or guild. It therefore cannot appropriately
327 capture the many variations in environmental associations of each species ([Graham et](#)
328 [al., 1996](#), [Jackson and Overpeck, 2000](#)) or address them with similar statistical power
329 as the *stacked-species* modeling approach, which fits many individual response curves.
330 Second, in the *direct community* modeling approach, the response variables are
331 strongly influenced by wide-ranging species, as they dominate grid cell records ([Jetz](#)
332 [and Rahbek, 2002](#)). This overwhelms the statistical signal attribute able to species
333 with more narrow ranges. The *stacked-species* modeling approach instead represents
334 the environmental associations of all species more uniformly and appropriately, which
335 in turn may contribute to their stronger predictive performance. Third, the
336 co-occurrence of species in communities and resulting patterns of absolute or relative
337 richness may be attributed more to regional factors or habitat filters than species
338 interactions ([Belmaker and Jetz, 2012](#), [Jetz and Fine, 2012](#)), especially at grains
339 coarser than those at which species interactions may play out ([Cooper and Belmaker,](#)
340 [2010](#)), as is likely the case in our study. As a consequence, environmental correlates
341 may not capture anything meaningful about species' co-occurrences, or at least may
342 not do so equally for all species. We note that the best predictive performance of the

direct community modeling approach was found for guilds that are known to face strong environmental constraints in their regional diversification or local coexistence; e.g. for obligate frugivores the requirement for fruits has obvious strong links with year-round high temperature and productivity (Kissling et al., 2009). Fourth, the ecological guild characterizations used here may yet be too broad to appropriately capture groups of interacting species suited for the one-group perspective of a *direct community* modeling approach. Our results did suggest better performance (i.e. lower absolute t values) for a few of the fine-grained guilds than the coarse-grain guilds.

Interestingly, we found relatively uniform guild prevalence along the gradient from low to high richness assemblages for most guilds, and especially so for high-richness assemblages. This might suggest there are “natural” restrictions on co-occurrence among environments, i.e. there is a maximum species richness attainable in a particular environment (environmental load). This may imply that the prevalence of different dietary guilds is maintained with a dynamic balance between the immigration of new species into the same or different guild and the extinction of species already present in a guild. This also may suggest that low richness assemblages in today's environment may not have yet reached a putative maximum environmental load.

Community-based conservation as a concept is relatively new and has burst into the center of the global conservation discourse in the past three decades. This concept espouses the integration of ecosystem management and human well-being and the requirements for linking protection and institutions at multiple levels and scales (Berkes, 2007, Gruber, 2011, Ruiz-Mallen and Corbera, 2013). Assemblage-level biodiversity patterns predicted by community-based modeling approaches, such as the *stacked-species* and *direct community* modeling approaches in this study, offer a

means to help guide conservation decisions for determining environmental suitability and analyzing diversity hotspots (Ferrier and Guisan, 2006, Guisan and Rahbek, 2011, Guisan et al., 2013). Moreover, incorporating community functional attributes, e.g. dietary guilds, with environmental responses in the modeling approaches reveals additional ecological insights. Besides, when species assemblages co-occur in range, habitat, and management tolerance, it is more able to use indicator species to monitor environmental changes, assess the efficacy of management, and provide warning signals for any impending ecological shifts, which further enhances the effectiveness of conservation projects (Carignan and Villard, 2002). In summary, our results highlight limitations for the use of *direct community* modeling approach for predicting geographic patterns in biodiversity. They support *stacked-species* modeling approach or methods combining *direct* and *stacked* elements, such as an emerging new generation of joint distribution models that combine co-occurrence with community signals (Pollock et al., 2014). The weak performance of *direct community* modeling approaches for current-day predictions calls into question their additional application for global biodiversity projections under global change. For some specific groups with apparent strong climatic constraints, the *direct community* modeling approach, including community richness and prevalence modeling, does appear to offer good predictive ability while requiring much fewer parameters than the *stacked-species* modeling approach. And at spatial grains much finer than those analyzed here (e.g. sub-kilometer), constraints arising from resource diversity or species interactions may favor methods using community attributes as response. But overall, our findings suggest caution against the utility of *direct community* modeling approach at the large spatial extents usually necessary to fully capture the environmental niches of groups of species.

Acknowledgements

We thank Jonathan Belmaker, Petr Keil, and Morgane Barbet-Massin for constructive suggestions of our work on modeling approaches and *Antoine Guisan*, Jui-Yi Hsieh and Ming-Fai Chow for helpful comments. We also thank two anonymous reviewers that made lots of insightful comments which improved the contents of this manuscript. We greatly appreciate financial support by multiple postdoc research fellowships funded by the Delta Electronics Foundation, Yu Chi-Chung Culture and Educational Foundation, and Taiwanese National Science Council (100IP999900127) to C.-Y. Ko and NSF grants DBI0960550 and DEB 1026764 and NASA Biodiversity Grant NNX11AP72G to W. Jetz.

References

- Ackerly, D.D., Cornwell, W.K., 2007. A trait-based approach to community assembly: partitioning of species trait values into within- and among-community components. *Ecol. Lett.* 10, 135-145.
- Algar, A.C., Kharouba, H.M., Young, E.R., Kerr, J.T., 2009. Predicting the future of species diversity: macroecological theory, climate change, and direct tests of alternative forecasting methods. *Ecography* 32, 22-33.
- Araújo, M.B., Luoto, M., 2007. The importance of biotic interactions for modelling species distributions under climate change. *Glob. Ecol. Biogeogr.* 16, 743-753.
- Araújo, M.B., New, M., 2007. Ensemble forecasting of species distributions. *Trends in Ecol. Evol.* 22, 42-47.

418 Baselga, A., Araújo, M.B., 2009. Individualistic vs community modelling of species
419 distributions under climate change. *Ecography* 32, 55-65.

420 Bellard, C., Bertelsmeier, C., Leadley, P., Thuiller, W., Courchamp, F., 2012. Impacts
421 of climate change on the future of biodiversity. *Ecol. Lett.* 15, 365-377.

422 Belmaker, J., Jetz, W., 2012. Regional pools and environmental controls of vertebrate
423 richness. *Am. Nat.* 179, 512-523.

424 Belmaker, J., Jetz, W., 2013. Spatial scaling of functional structure in bird and
425 mammal assemblages. *Am. Nat.* 181, 464-478.

426 Berkes, F., 2007. Community-based conservation in a globalized world. *PNAS* 104,
427 15188-15193.

428 Bonthoux, S., Baswlg, A., Balent, G., 2013. Assessing community-level and
429 single-species models predictions of species distributions and assemblage
430 composition after 25 years of land cover change. *PLoS One* 8, e54179.

431 Brown, J.H., 1981. Two decades of homage to Santa Rosalia: toward a general theory
432 of diversity. *Am. Zool.* 21, 877-888.

433 Calabrese, J.M., Certain, G., Kraan, C., Dormann, C.F., 2014. Stacking species
434 distribution models and adjusting bias by linking them to macroecological
435 models. *Glob. Ecol. Biogeogr.* 23, 99-112.

436 Carignan, V., Villard, M.-A., 2002. Selecting indicator species to monitor ecological
437 integrity: review. *Environmental Monitoring and Assessment* 78, 45-61.

438 Chapman, D.S., Purse, B.V., 2011. Community versus single-species distribution
439 models for British plants. *J. Biogeogr.* 38, 1524-1535.

440 Chown, S.L., le Roux, P.C., Ramaswiela, T., Kalwijn, J.M., Shaw, J.D., McGeoch,
441 M.A., 2012. Climate change and elevational diversity capacity: do weedy species
442 take up the slack? *Biol. Lett.* 9, 20120806.

443 Cooper, N., Belmaker, J., 2010. Opinion: habitat data resolution and the detection of
 444 species interactions. *Front. Biogeogr.* 2, 46.

445 D'Amen, M., Dubuis, A., Fernandes, R.F., Pottier, J., Pellissier, L., Guisan, A., 2015.
 446 Using species richness and functional traits predictions to constrain assemblage
 447 predictions from stacked species distribution models. *J. Biogeogr.* 42, 1255-1266.

448 Dawson, T.P., Jackson, S.T., House, J.I., Prentice, I.C., Mace, G.M., 2011. Beyond
 449 predictions: biodiversity conservation in a changing climate. *Science* 332, 53-58.

450 Distler, T., Schuetz, J.G., Velasquez-Tibata, J., Langham, G.M., 2015. Stacked species
 451 distribution models and macroecological models provide congruent projections
 452 of avian species richness under climate change. *J. Biogeogr.* 42, 976-988.

453 Dubuis, A., Pottier, J., Rion, V., Pellissier, L., Theurillat, J.-P., Guisan, A., 2011.
 454 Predicting spatial patterns of plant species richness: a comparison of direct
 455 macroecological and species stacking modelling approaches. *Divers. Distrib.* 17,
 456 1122-1131.

457 Dubuis, A., Rossier, L., Pottier, J., Pellissier, L., Vittoz, P., Guisan, A., 2013.
 458 Predicting current and future spatial community patterns of plant functional
 459 traits. *Ecography* 36, 1158-1168.

460 Elith, J., Graham, C.H., Anderson, R.P., Dudík, M., Ferrier, S., Guisan, A., Hijmans,
 461 R.J., Huettmann, F., Leathwick, J.R., Lehmann, A., Li, J., Lohmann, L.G.,
 462 Loiselle, B.A., Manion, G., Moritz, C., Nakamura, M., Nakazawa, Y., Overton,
 463 J. McC. M., Peterson, A.T., Phillips, S.J., Richardson, K., Scachetti-Pereira, R.,
 464 Schapire, R.E., Soberón, J., Williams, S., Wisz, M.S., Zimmermann, N.E., 2006.
 465 Novel methods improve prediction of species' distributions from occurrence
 466 data. *Ecography* 29, 129-151.

467 Elith, J., Leathwick, J.R., 2007. Predicting species distribution from museum and

468 herbarium records using multiresponse models fitted with multivariate adaptive
 469 regression splines. *Divers. Distrib.* 13, 265-275.

470 Ferrier, S., Guisan, A., 2006. Spatial modelling of biodiversity at the community
 471 level. *J. Appl. Ecol.* 43, 393-404.

472 Field, R., Hawkins, B.A., Cornell, H.V., Currie, D.J., Diniz-Filho, J.A.F., Guégan,
 473 J.-F., Kaufman, D.M., Kerr, J.T., Mittelbach, G.G., Oberdorff, T., O'Brien, E.M.,
 474 Turner, J.R.G., 2009. Spatial species-richness gradients across scales: a
 475 meta-analysis. *J. Biogeogr.* 36, 132-147.

476 Graham, R.W., Lundelius Jr., E.L., Graham, M.A., Schroeder, E.K., Toomey III, R.S.,
 477 Anderson, E., Barnosky, A.D., Burns, J.A., Churcher, C.S., Grayson, D.K.,
 478 Guthrie, R.D., Harington, C.R., Jefferson, G.T., Martin, L.D., McDonald, H.G.,
 479 Morlan, R.E., Semken Jr., J.A., Webb, S.D., Werdelin, L., Wilson, M.C., 1996.
 480 Spatial response of mammals to late Quaternary environmental fluctuations.
 481 *Science* 272, 1601-1606.

482 Gruber, J.S., 2011. Perspectives of effective and sustainable community-based natural
 483 resource management: an application of Q methodology to forest projects.
 484 *Conserv. Soc.* 9, 159-171.

485 Guillera-Aroita, G., Lahoz-Monfort, J.J., Elith, J., Gordon, A., Kujala, H., Lentini,
 486 P.E., McCarthy, M.A., Tingley, R., Wintle, B.A., 2015. Is my species distribution
 487 model fit for purpose? matching data and models to applications. *Glob. Ecol.*
 488 *Biogeogr.* 24, 276-292.

489 Guisan, A., Rahbek, C., 2011. SESAM – a new framework integrating
 490 macroecological and species distribution models for predicting spatio-temporal
 491 patterns of species assemblages. *J. Biogeogr.* 38, 1433-1444.

492 Guisan, A., Tingley, R., Baumgartner, J.B., Naujokaitis-Lewis, I., Sutcliffe, P.R.,

493 Tulloch, A.I.T., Regan, T.J., Brotons, L., McDonald-Madden, E.,
 494 Mantyka-Pringle, C., Martin, T.G., Rhodes, J.R., Maggini, R., Setterfield, S.A.,
 495 Elith, J., Schwartz, M.W., Wintle, B.A., Broennimann, O., Austin, M., Ferrier, S.,
 496 Kearney, M.R., Possingham, H.P., Buckley, Y.M., 2013. Predicting species
 497 distributions for conservation decisions. *Ecol. Lett.* 16, 1424-1435.
 498 Guisan, A., Zimmermann, N.E., 2000. Predictive habitat distribution models in
 499 ecology. *Ecol. Model.* 135, 147-186.
 500 Heikkinen, R.K., Luoto, M., Virkkala, R., Pearson, R.G., Korber, J.-H., 2007. Biotic
 501 interactions improve prediction of boreal bird distributions at macro-scales.
 502 *Glob. Ecol. Biogeogr.* 16, 754-763.
 503 Hurlbert, A.H., Jetz, W., 2007. Species richness, hotspots, and the scale dependence of
 504 range maps in ecology and conservation. *PNAS* 104, 13384-13389.
 505 Jackson, S.T., Overpeck, J.T., 2000. Responses of plant populations and communities
 506 to environmental changes of the late Quaternary. *Paleobiology* 26, 194-220.
 507 Jetz, W., Rahbek, C., 2002. Geographic range size and determinants of avian species
 508 richness. *Science* 297, 1548-1551.
 509 Jetz, W., Wilcove, D.S., Dobson, A.P., 2007. Projected impacts of climate and
 510 land-use change on the global diversity of birds. *PLoS Biol.* 5, 1211-1219.
 511 Jetz, W., Fine, P.V., 2012. Global gradients in vertebrate diversity predicted by
 512 historical area-productivity dynamics and contemporary environment. *PLoS*
 513 *Biol.* 10, e1001292.
 514 Jetz, W., Thomas, G.H., Joy, J.B., Hartmann, K., Mooers, A.O., 2012. The global
 515 diversity of birds in space and time. *Nature* 491, 444-448.
 516 Kardol, P., Cregger, M.A., Campany, C.E., Classen, A.T., 2010. Soil ecosystem
 517 functioning under climate change: plant species and community effects. *Ecology*

518 91, 767-778.

519 Kissling, W.D., Rahbek, C., Böhning-Gaese, K., 2007. Food plant diversity as
520 broad-scale determinant of avian frugivore richness. *Proc. R. Soc. Lond. B Biol.*
521 *Sci.* 274, 799-808.

522 Kissling, W. D., Böhning-Gaese, K., Jetz, W., 2009. The global distribution of
523 frugivory in birds. *Glob. Ecol. Biogeogr.* 18, 150-162.

524 Kissling, W.D., Böhning-Gaese, K., Jetz, W., 2010. Woody plants and the prediction
525 of climate-change impacts on bird diversity. *Phil. Trans. R. Soc. B.* 365,
526 2035-2045.

527 Kissling, W.D., Sekercioglu, C.H., Jetz, W., 2012. Bird dietary guild richness across
528 latitudes, environments and biogeographic regions. *Glob. Ecol. Biogeogr.* 21,
529 328-340.

530 Ko, C.-Y., Schmitz, O.J., Barbet-Massin, M., Jetz, W., 2014. Dietary guild
531 composition and disaggregation of avian assemblages under climate change.
532 *Glob. Change Biol.* 20, 790-802.

533 Leathwick, J.R., Elith, J., Hastie, T., 2006. Comparative performance of generalized
534 additive models and multivariate adaptive regression splines for statistical
535 modelling of species distributions. *Ecol. Model.* 199, 188-196.

536 Lemoine, N., Böhning-Gaese, K., 2003. Potential impact of global climate change on
537 species richness of long-distance migrants. *Conserv. Biol.* 17, 577-586.

538 Marmion, M., Parviainen, M., Luoto, M., Heikkinen, R.K., Thuiller, W., 2009.
539 Evaluation of consensus methods in predictive species distribution modelling.
540 *Divers. Distrib.* 15, 59-69.

541 McPherson, J.M., Jetz, W., Rogers, D.J., 2004. The effects of species' range sizes on
542 the accuracy of distribution models: ecological phenomenon or statistical

543 artefact? J. Appl. Ecol. 41, 811-823.

544 Mitchell, T.D., Jones, P.D., 2005. An improved method of constructing a database of
545 monthly climate observations and associated high-resolution grids. International
546 J. Climatol. 25, 693-712.

547 Mooney, H., Larigauderie, A., Cesario, M., Elmquist, T., Hoegh-Guldberg, O.,
548 Lavorel, S., Mace, G.M., Palmer, M., Scholes, R., Yahara, T., 2009. Biodiversity,
549 climate change, and ecosystem services. Curr. Opin. Env. Sust. 1, 46-54.

550 Pearson, R.G., Raxworthy, C.J., Nakamura, M., Peterson, A.T., 2006. Predicting
551 species distributions from small numbers of occurrence records: a test case using
552 cryptic geckos in Madagascar. J. Biogeogr. 34, 102-117.

553 Peterson, A.T., Papeş, M., Eaton, M., 2006. Transferability and model evaluation in
554 ecological niche modeling: a comparison of GARP and Maxent. Ecography 30,
555 550-560.

556 Peterson, A.T., Shaw, J.J., 2003. *Lutzomyia* vectors for cutaneous leishmaniasis in
557 southern Brazil: ecological niche models, predicted geographic distributions, and
558 climate change effects. Int. J. Parasitol. 33, 919-931.

559 Piñeiro, G., Perelman, S., Guerschman, J.P., Paruelo, J., 2008. How to evaluate
560 models: observed vs. predicted or predicted vs. observed? Ecol. Model. 216,
561 316-322.

562 Pollock, L.J., Tingley, R., Morris, W.K., Golding, N., O'Hara, R.B., Parris, K.M.,
563 Vesk, P.A., McCarthy, M.A., 2014. Understanding co-occurrence by modelling
564 species simultaneously with a Joint Species Distribution Model (JSDM).
565 Methods Ecol. Evol. 5, 397-406.

566 R Development Core Team, 2011. R: A Language and Environment for Statistical
567 Computing. R Foundation for Statistical Computing, Vienna, Austria.

568 Ruiz-Mallen, I., Corbera, E., 2013. Community-based conservation and traditional
569 ecological knowledge: implications for social-ecological resilience. *Ecol. Soc.*
570 18, 12.

571 Schmitz, O.J., Post, E., Burns, C.E., Johnston, K.M., 2003. Ecosystem responses to
572 global climate change: moving beyond color-mapping. *BioScience* 53,
573 1199-1205.

574 Sommer, J.H., Kreft, H., Kier, G., Jetz, W., Mutke, J., Barthlott, W., 2010. Projected
575 impacts of climate change on regional capacities for global plant species
576 richness. *Proc. R. Soc. Lond. B Biol. Sci.* 277, 2271-2280.

577 Steinmann, K., Linder, H.P., Zimmermann, N.E., 2009. Modelling plant species
578 richness using functional groups. *Ecol. Model.* 220, 962-967.

579 Thuiller, W., Araújo, M.B., Lavorel, S., 2003. Generalized models vs. classification
580 tree analysis: predicting spatial distributions of plant species at different scales. *J.*
581 *Veg. Sci.* 14, 669-680.

582 Thuiller, W., Lafourcade, B., Engler, R., Araújo, M.B., 2009. BIOMOD - a platform
583 for ensemble forecasting of species distributions. *Ecography* 32, 369-373.

584 VanDerWal, J., Shoo, L.P., Graham, C., Williams, S.E., 2009. Selecting
585 pseudo-absence data for presence-only distribution modeling: how far should
586 you stray from what you know? *Ecol. Model.* 220, 589-594.

587 Venevskaya, I., Venevsky, S., Thomas, C.D., 2013. Projected latitudinal and regional
588 changes in vascular plant diversity through climate change: short-term gains and
589 longer-term losses. *Biodivers. Conserv.* 22, 1467-1483.

590 Weiher, E., Keddy, P.A., 1999. *Ecological assembly rules : perspectives, advances,*
591 *retreats.* Cambridge University Press, Cambridge.

592 Wilman, H., Belmaker, J., Simpson, J., de la Rosa, C., Rivadeneira, M.M., Jetz, W.,

593 2014. EltonTraits 1.0: Species-level foraging attributes of the world's birds and
594 mammals. Ecology 95, 2027.

Table 1

Variation in the observed and predicted prevalence (proportional richness) of avian dietary guilds, modeled with *stacked-species* vs. *direct community* modeling approaches. Values in the table are t statistic (t) for the difference from a slope of 1 and root mean squared deviation (RMSD) from linear regression of predicted vs. observed prevalence, calculated from the predictions by the stacked-species, community richness and prevalence modeling approaches.

Dietary guild		Stacked-species approach		Direct community approach			
		t	RMSD	richness model		prevalence model	
				t	RMSD	t	RMSD
Coarse	Primary consumers	13.8	0.015	-391.1	0.04	-424.6	0.04
	High-level consumers	7.6	0.017	-320.2	0.046	-318.3	0.047
	Mixed consumers	-1.5	0.008	-226	0.02	-222.1	0.02
Fine	Frugivores	25.4	0.004	-37	0.014	-36.1	0.014
	Nectarivores	-18.5	0.003	-89.4	0.013	-99.8	0.013
	Herbivores	12.7	0.013	-139.4	0.029	-92.5	0.026
	Granivores	37.3	0.012	-150	0.029	-128.7	0.028
	Insectivores	4.2	0.025	-145.4	0.048	-113.2	0.046
	Carnivores	45.6	0.02	-162.3	0.044	-172.3	0.045
	Scavengers	-54.8	0.008	-169.3	0.011	-134.2	0.01
	Omnivores	23.6	0.015	-122	0.042	-131.4	0.041

Fig. 1. Relationships between observed and predicted present (2000) relative richness/prevalence (proportions of total) of three coarse dietary guilds, i.e. primary consumers (a), high-level consumers (b) and mixed consumers (c), using *stacked-species* and *direct community* modeling approaches for the global bird fauna (8,472 species; 11,079 110 x 110 km grid cells worldwide). Red dashed lines represent a 1:1 relationship.

Fig. 2. Relationships between observed and predicted present (2000) relative richness/prevalence (proportions of total) of eight fine dietary guilds, i.e. frugivores (a), nectarivores (b), herbivores (c), granivores (d), insectivores (e), carnivores (f), scavengers (g) and omnivores (h), using *stacked-species* and *direct community* modeling approaches for the global bird fauna (8,472 species; 11,079 110 x 110 km grid cells worldwide). Red dashed lines represent a 1:1 relationship.

Fig. 3. Scatterplots of R-squared values from linear regression of predicted vs. observed prevalence for three coarse and eight fine dietary guilds using *stacked-species* and *direct community* modeling approaches. Dashed lines represent a 1:1 relationship. Abbreviations of the dietary guilds: P, primary consumers; H, high-level consumers; M, mixed consumers; Fru, Frugivores; Nec, Nectarivores; Her, Herbivores; Gra, Granivores; Ins, Insectivores; Car, Carnivores; Sca, Scavengers; Omn, Omnivores.

Fig. 4. Relationships between current-day observed total richness and predicted guild prevalence from different modeling approaches across three coarse dietary guilds. Each point is the species richness and the proportion of a guild in a grid cell.

Fig. 5. Relationships between current-day observed total richness and predicted guild prevalence from different modeling approaches across eight fine dietary guilds. Each point is the species richness and the proportion of a guild in a grid cell. Both in the observations and predictions, frugivores show increasing guild prevalence with the increasing total richness while most fine dietary guilds maintain prevalence across the total richness gradient especially under high species richness.