DECISIONS AND ACTIONS OF DISTRACTED DRIVERS AT THE ONSET OF YELLOW LIGHTS

Abstract

Driving on an approach to a signalized intersection while distracted is relatively risky, as potential vehicular conflicts and resulting angle collisions tend to be relatively more severe compared to other locations. Given the prevalence and importance of this particular scenario, the objective of this study was to examine the decisions and actions of distracted drivers during the onset of yellow lights. Driving simulator data were obtained from a sample of 69 drivers under baseline and handheld cell phone conditions at the University of Iowa - National Advanced Driving Simulator. Explanatory variables included age, gender, cell phone use, distance to stop-line, and speed. Although there is extensive research on drivers' responses to yellow traffic signals, the examinations have been conducted from a traditional regression-based approach, which do not necessary provide the underlying relations and patterns among the sampled data. In this paper, we exploit the benefits of both classical statistical inference and data mining techniques to identify the *a priori* relationships among main effects, non-linearities, and interaction effects. Results suggest that the probability of yellow light running increases with the increase in driving speed at the onset of yellow. Both young (18-25 years) and middle-aged (30-45 years) drivers reveal reduced propensity for yellow light running whilst distracted across the entire speed range, exhibiting possible risk compensation during this critical driving situation. The propensity for yellow light running for both distracted male and female older (50-60 years) drivers is significantly higher. Driver experience captured by age interacts with distraction, resulting in their combined effect having slower physiological response and being distracted particularly risky.

Key words: Distracted driving; Mobile phone; Motion-based driving simulator; Risk compensation; Yellow light; Driver behavior

1. Introduction

1.1 Distracted Driving

Driver distraction is a significant and growing road safety issue worldwide. Statistics reveal that about 3,328 people were killed and an additional 421,000 people were injured in distracted driving related motor vehicle crashes on U.S. roadways in 2012 (NHTSA, 2014). While cell phone use has become ubiquitous in modern society, road traffic crashes related to cell phone distractions have been on the rise (WHO, 2011). In 2012, cell phone distraction alone was associated with 415 fatalities and another 28,000 injuries on U.S. roads (NHTSA, 2014). Redelmeier and Tibshirani (1997) indicated that distraction due to conversing on a cell phone while driving increased the crash risk by as much as four folds. In 2010, the traffic police of Queensland, Australia fined about 30,780 drivers for talking or texting on mobile phones while driving, indicating the spread of cell phone usage (Ironsides, 2011). Making matters worse, distracted driving tends to be more prevalent among younger and less experienced drivers. An Australian study reported that among the 2400 driver distraction related incidents in New South Wales, young drivers had the highest frequency of cell phone use-related injurious crashes (Lam, 2002).

Prior research has documented a variety of performance measures that have been impacted by the distracting effects of a cell phone. Burns et al. (2002) reported that speed control and reaction times of drivers were more influenced by a cell phone use than by having a blood alcohol level at the limit of 8%. A recent study reported that cognitive distraction significantly impairs the reaction time of young drivers in response to a traffic event that originates within the driver's peripheral vision (Haque & Washington, 2014a). Cell phone distraction has also been reported influencing drivers to have higher variation in accelerator pedal position, drive more slowly with greater speed variation, and report a higher workload (e.g., Rakauskas, Gugerty, & Ward, 2004). Tornos and Bolling (2006) reported a risk compensation behavior of distracted drivers where drivers tend to reduce their speed while talking on a phone. Dula et al. (2011) showed that the percent of time spent speeding and the number of centre line crossings were significantly higher among drivers engaged in different types of conversation in comparison to no conversation.

Hancock et al. (2003) investigated the stopping decision of a group of cell phone distracted drivers using test track facilities where participants were instructed to perform a quick stop before reaching the stop line of an intersection upon the onset of a red light. They found that the non-response to a red light increased by 15% for drivers distracted by a cell phone visual-manual task consisting of looking down at a dial pad and confirming a number presented was the same as the number presented earlier. Using a driving simulator, Beede and Kass (2006) found that cell phone distracted drivers took one-third of a second longer before starting from a stop-sign indicating a slower response of distracted drivers. Consiglio et al. (2003) examined the braking performance of distracted drivers upon the activation of a red brake lamp set at a laboratory station and found that both hands-free and hand-held cellular phone conversation significantly increased the reaction time to braking events initiated by the lamp. Haque and Washington (2014b) examined the braking behavior of distracted drivers in response to a pedestrian entering a zebra crossing, and reported that drivers distracted either by handheld or hands-free phone conversation are associated with aggressive braking compared to non-distracted drivers, revealing perhaps an element of risk compensation. Strayer and Johnston (2001) examined the effects of cellular phone conversation on driving performances using a desktop simulator study, where participants performed a pursuit tracking task with a joystick and responded to flashing signals on a computer display. The study reported that drivers distracted by hand-held or hands-free conversation were two times

less likely to detect simulated traffic signals and exhibited slower reactions to those signals that were detected.

In summary, using a cell phone while driving appears to influence many common driving behaviors including a deterioration of speed control, reduction in speed, failure to maintain appropriate headway, increase in lane position variation, reduction of peripheral eye scanning, decline in braking performances, and impairment in the perception of relevant stimuli (e.g., Regan, Young, & Lee, 2009). In particular, the slower reaction and impaired braking performances of distracted drivers might affect their safety at the onset of yellow lights at signalized intersection. However, there is little research on how distracted drivers perform at the onset of yellow lights.

1.2 Driver Behavior at the Onset of Yellow Light

Yellow light or phase changing period, the time interval when drivers need to decide on stopping or proceeding through an intersection, has been identified as a critical interval at signalized intersections (e.g., Elmitiny, Yan, Radwan, Russo, & Nashar, 2010; Papaioannou, 2007). An improper decision at the onset of a yellow or amber light might cause red light running or abrupt stopping at intersections. It was estimated that red light running is associated with about 260,000 crashes on U.S. roads each year, of which about 750 are fatal (Retting, Ulmer, & Williams, 1999). Earlier research on the driver's stop/go decisions at the onset of yellow lights has mainly been focused on developing a model to estimate the propensity of yellow running as a function of driving speed, distance from the stop line, and drivers' demographics like age and gender. Papaioannou (2007) reported that drivers who have an approach speed higher than the posted speed limit are more likely to be caught in a dilemma zone that might force them to make unsafe decisions at the onset of yellow light. Elmitiny et al. (2010) analyzed video-based data of a high-speed signalized intersection and found that red light violations and drivers' stop/go decisions during the yellow light are significantly associated with their distance from the stop line, operating speed, and position in the traffic flow. Using a driving simulator, Caird et al. (2007) investigated younger and older drivers' behavior at the onset of yellow light at signalized intersections and found that all drivers are less likely to run through the yellow light when their time to stop line is higher. Ohlhauser et al. (2011) investigated the effects of distraction at the onset of yellow light across various age cohorts. They reported that compared to middle-aged drivers, novice drivers are more likely to proceed through the intersection when distracted by a handheld phone conversation.

1.3 Research Objective

Drivers' stop/go decisions have previously been modeled in relation to driver demographics such as age and gender, and traffic parameters such as speed and distance from the stop line. It is of interest to examine whether distraction will impact a driver's response to the onset of a yellow interval while approaching a signalized intersection. In particular, it is worthy to investigate how a cell phone distraction interacts with age, gender differences and different traffic parameters in relation to the stop/go decisions of drivers at the onset of yellow light. While it might be difficult to expose distracted drivers at yellow light encounters in real world settings repeatedly, a driving simulator could be fully utilized to include such scenarios. The objective of this study is to examine the decisions of distracted drivers at the onset of yellow lights by exposing a group of distracted drivers on a series of signalized intersections using the National Advanced Driving Simulator at the University of Iowa.

2. Experiment Details

2.1 Participants

The data analyzed in this study were gathered from two separate studies conducted at the University of Iowa: a wireless urban arterial study and a novice driver study (Mazzae, Goodman, Garrott, & Ranney, 2005; Ranney et al., 2005). The wireless urban arterial study on three adult groups was conducted in the summer of 2004, and the novice driver study on 16-17 years old male drivers was conducted between May and October of 2006. Details of recruitment, screening, and compensation are available in Marshall et al. (2010).

In total, there were 69 drivers including 49 adults and 20 novice drivers. A valid driver's license was the main criterion for adult drivers and at least 4 weeks but no more than 8 weeks of licensure was the criterion for novice drivers. While the 20 male novice drivers of 16-17 years old formed a separate group, the 49 adult participants were divided into the following three age groups: younger (18-25 years old including 9 male and 9 female), middle-aged (30-45 years old including 9 male and 8 female), and older (50-60 years old including 8 male and 6 female). The average driving experiences for novice, younger, middle, and older drivers were 0.1, 5.5, 19.8, and 36.7 years, respectively.

2.2 Driving Simulator

Both studies were conducted in the high fidelity NADS-1 simulator located at the University of Iowa. It is known to be the most advanced of its kind, consisting of an entire car housed inside a 24-foot dome providing 360-degree high resolution field-of-view to drivers. Road images and interactive traffic are updated at 60 Hz on eight LCD projectors to provide a photorealistic virtual environment. It is comprised of a 13 degree-of-freedom motion base to accurately reproduce motion cues for sustained acceleration and braking maneuvers, movement across multiple lanes of traffic, and interaction with varying road surfaces. Driving performance data such as lane position, speed, acceleration, and braking are recorded at rates up to 240 Hz.

2.3 Procedure

Prior to the experiment, each participant, including the parent/guardian of novice drivers, completed an informed consent. After providing an overview of the study drives, participants were briefed about how to use the simulator and the cell phone apparatus. Before participating in the experimental drive, each participant performed a practice drive to be familiar with the driving simulator.

The traffic scenarios and distracting tasks were the same for both studies, and details are provided in Ranney et al. (2005) and Marshall et al. (2010). While the aim of Ranney et al. (2005) was to examine the effects of wireless phone use on driver performance for adults, the aim of Marshall et al. (2010) was to examine the effects of the same distracting task on novice drivers (16-17 years old). Although these original studies did not focus specifically on the risk of yellow light running, they used the same simulator protocol and contained detailed scenarios for traffic light change on a series of signalized intersections.

In summary, the driving route on the NADS-1 simulator contained three segments (approximately 15 minute each) with each segment containing both an urban and rural area. Five controlled intersections were included in each segment (or 15 total). The analyses in this paper focused on the urban areas where only two signalized intersections existed per segment (or six total). Programs were scripted so that the traffic light turned from green to yellow

1 when the host vehicle was at 3.0 seconds from the stop line in one intersection and at 3.75
2 seconds from the stop line in another. The yellow phase was also 3.0 and 3.75 seconds
3 respectively to generate a “dilemma zone” situation. The time to activate the yellow lights
4 was counterbalanced across all segments. The road consisted of four lanes (12 feet each) with
5 two lanes in each direction. There was ambient traffic traveling in the same direction as the
6 participant (northbound) at level of service A (free flowing) and there was no traffic in the
7 opposite direction. The posted speed limit on the approach was 45 mph. At the 4-way
8 signalized intersection of interest, there was no other traffic present during the dilemma zone.

9 Among the three driving segments, one segment was used as a baseline condition, i.e.,
10 without any distraction task and the other two segments involved participants talking with the
11 experimenter using a handheld phone interface. That is, each participant drove through two
12 intersections without any conversation task and four intersections with a phone conversation
13 task. This included placing a call and answering a call. Driving events such as change of
14 traffic light from green to yellow occurred while participants were in the conversation phase
15 of the call. The order of the driving segments for baseline and phone conversation tasks was
16 counterbalanced across participants. It should be noted that proceeding through a yellow light
17 does not necessarily comprise a surprise event given that there is no possibility of the driver
18 colliding into another vehicle. Hence, it is reasonable to have repeated events per participants.
19 Further, running a yellow light does not necessarily indicate an undesirable outcome nor is it
20 considered a moving violation in the US.

21 A modified version of the Baddeley working memory span task was used (Baddeley,
22 Logie, Nimmo-Smith, & Brereton, 1985; Donmez, Boyle, & Lee, 2006) to engage
23 participants in conversation while driving. The working memory span task requires
24 simultaneous storage and processing of information, and thus distracts the drivers by
25 increasing the cognitive loads. In this task, participants were required to listen to sentences
26 and determine whether they made sense. The experiment was designed to ensure that the
27 conversation lasted for the duration of the driving events.

28 29 **2.5 Data Collection**

30
31 The driver’s decision to proceed through or stop at the onset of yellow light was extracted as
32 a binary outcome from simulator data and included as the dependent variable. Explanatory
33 variables included driver’s age group, gender, phone condition and speed at the onset of
34 yellow. Phone condition variable had two categories: baseline and handheld conversation.

35 Traffic lights in the driving simulator were programmed using a time to stop line
36 (TSL) variable, which represents the time required for a participant’s vehicle to reach the stop
37 line based on their speed and distance from the stop line at the onset of yellow light. In
38 particular, the traffic lights were programmed to change from green to yellow when the
39 participants were 3 or 3.75 seconds from the stop line. To account for any variability, TSL at
40 the onset of yellow light was calculated from the dataset and included as a continuous
41 variable for analysis. The minimum and maximum TSL were 2.47 and 3.80 seconds
42 respectively. Mean TSL was 3.27 seconds with a standard deviation of 0.37.

43 In total, there were 414 observations for 69 drivers encountering a traffic light change
44 in six urban intersections along three road segments. Driving performance data such as speed
45 of the driven car and distance from the stop line at the onset of yellow in five intersections
46 were not correctly captured by the driving simulator due to a data recording error. As a result,
47 there were 409 total observations for 69 drivers representing an unbalanced panel data with
48 minimum 2 and maximum 6 observations per driver. Summary statistics of the variables for
49 the decision tree and subsequent logistic regression model are presented in table 1. There
50 were 119, 108, 102 and 80 encounters with the traffic light change at signalized intersections

for novice, young, middle-aged and older drivers, respectively. Among them 136 encounters with traffic light change happened in baseline or no phone conversation driving, while the rest (273) encounters happened in handheld phone conversation driving. The mean driving speed at the onset of yellow light was 42.5 mph with a standard deviation of 5.2. Among the total 409 observations at signalized intersections, drivers decided to proceed through a yellow light in 153 encounters with the traffic light change.

3. Statistical Modeling

A repeated measures logistic regression model in the form of Generalized Estimation Equations (GEE) was applied to account for multiple observations across individuals. GEEs are an extension of generalized linear models to analyze correlated data, where the correlation is the result of repeated observations of the same driver at multiple points in time. GEEs are quite flexible in that they can accommodate non-normal variables and non-linear relationships well. In GEEs, the marginal expectation of the dependent variable is specified as a known linear function of covariates, assuming that the variance is a known function of the mean. In addition, GEEs specify a ‘working’ correlation matrix for the observations of each driver (Liang & Zeger, 1986).

Suppose, the binary outcome of driver’s decision to proceed through or stop at the onset of yellow light for driver i at intersection j is Y_{ij} ($i = 1, 2, \dots, k; j = 1, 2, \dots, n_i$). In the logistic regression model, the marginal expectation of, $E(Y_{ij}) = \mu_{ij}$, satisfies $\log it(\mu_{ij}) = \mathbf{X}'_{ij}\boldsymbol{\beta}$, where $\mathbf{X}'_{ij} = (X_{ij1}, \dots, X_{ijp})'$ denote a $P \times 1$ vector of explanatory variables, and $\boldsymbol{\beta}$ is a vector of estimable regression parameters. Then the probability of yellow light running of driver i could be expressed as:

$$\Pr(Y_{ij} = 1) = \frac{\exp(\mathbf{X}'_{ij}\boldsymbol{\beta})}{1 + \exp(\mathbf{X}'_{ij}\boldsymbol{\beta})} \quad (1)$$

Suppose, V_i is an estimator of the covariance matrix of Y_i , then $\boldsymbol{\beta}$ can be estimated by solving the GEEs as follows:

$$S(\boldsymbol{\beta}) = \sum_{i=1}^k \left(\frac{d\mu_i}{d\boldsymbol{\beta}} \right)' V_i^{-1} (Y_i - \mu_i(\boldsymbol{\beta})) = 0 \quad (2)$$

The estimator of the covariance matrix is specified as $V_i = \phi A_i^{1/2} R_i(\rho) A_i^{1/2}$, where A_i is a $n_i \times n_i$ diagonal matrix with $v(\mu_{ij})$ as the j^{th} diagonal element. V_i can be different from one subject to another, but it is common to specify the same form of V_i for all subjects. $R_i(\rho)$ is a $n_i \times n_i$ working correlation matrix that is fully specified by the vector parameters ρ . An exchangeable working correlation that makes constant correlations between any two observations within a subject is specified as:

$$\text{Corr}(Y_{ij}, Y_{il}) = \begin{cases} 1 & j = l \\ \rho & j \neq l \end{cases} \quad \text{e.g.,} \quad R_{6 \times 6} = \begin{bmatrix} 1 & \rho & \rho & \rho & \rho & \rho \\ \rho & 1 & \rho & \rho & \rho & \rho \\ \rho & \rho & 1 & \rho & \rho & \rho \\ \rho & \rho & \rho & 1 & \rho & \rho \\ \rho & \rho & \rho & \rho & 1 & \rho \\ \rho & \rho & \rho & \rho & \rho & 1 \end{bmatrix} \quad (3)$$

Detailed expressions for estimating ρ 's are available in Liang and Zeger (1986). In addition, robust variance estimates have been applied because they are agnostic about the nature of the interdependence in the data (Zorn, 2006). That is, the robust variance estimates do not depend on whether the conditional correlation among observations is positive or negative and thus provides better estimates of the parameters of interest.

Selecting the best subset of explanatory variables that includes main effects, non-linearities, and interactions effects is often challenging, especially when little prior experience guides a specification. This is because traditional approaches require the analyst to specify a priori second- and higher-order interactions, non-additive behaviour, and non-linearities of the main effect variables prior to running stepwise or best subsets procedures. In many explanatory research settings, the number of the potential combinations of variables and their higher-order interactions grow geometrically with the number of ordinal-scale variable and exponentially with nominal-scale variables. Therefore, it is difficult to decide which interactions to test and which to omit.

To circumvent this problem, an iterative combination of decision tree and logistic regression analyses was applied in this study. Decision or classification and regression trees are nonparametric methods; they are exploratory in nature and are helpful for uncovering possible relationships by selecting an appropriate explanatory variable in each stage to produce a maximum reduction in variability of the dependent variable; however, decision trees suffer from type I errors and lack the ability to make inferences. Despite these limitations, the results of a decision tree can be utilized to provide *a priori* knowledge for the logistic regression model by identifying interactions indicated by the tree branches. Therefore, this combined modeling technique allows exploiting the best of both qualities: possible higher order interactions by decision trees and inferences by the logistic regression model (Washington, 2000). The general procedures for including higher order interactions can be summarized in the following steps:

1. Estimate a decision tree to reveal patterns and relationships of the sampled data.
2. Convert the interactions identified by the decision tree into indicator variables.
3. Estimate a logistic regression model with both main effect variables and interaction variables.
4. Retain significant interactive variables, prune the tree and re-identify possible interactions
5. Iterate steps 3-4 until a theoretically and practically appealing justifiable specification of the logistic regression model is obtained.

In summary, the combined approach provides researchers *a priori* knowledge to select interaction terms and is helpful for conducting exploratory research and developing statistically and theoretically defensible model specifications.

In order to interpret the effect of coefficient estimates, the exponents of the parameter estimates (i.e. $\exp(\beta)$) were calculated to obtain the odds ratios (ORs) which provide the magnitude of the association between the factor of interest and the probability of running a yellow light. ORs greater than 1 indicate increased probability whereas ORs less than 1 indicate decreased probability. For categorical variables, $\exp(\beta_a - \beta_b)$ is used to represent the odds ratio between two categories. All parameter estimates were mainly assessed at 5% significance level, and ORs are reported with the corresponding 95% confidence intervals (CI).

4. Results

A decision tree was constructed from available data to classify discrete outcomes of driver's stop/go decision at the onset of yellow indications at signalized intersections. Variables like driver's age group, gender, distance to stop line, speed, and time to stop line (TSL) were offered as possible explanatory variables in the decision tree. Figure 1 illustrates the decision tree diagram for the stop/go decision model. Nodes were determined selecting the option that offered the highest information gain. The decision tree was constructed using 10-fold stratified cross-validation, which divided the data into ten unique partitions, which were each used in turn to test the decision tree. Thus, on each cycle nine-tenths of the data was being used to train the decision tree. The decision tree correctly classified 60% of instances, using 15 leaves for a total tree size of 27 nodes. Driver's age represents the highest information gain and is therefore at the top of the tree (as shown in Figure 1). Higher order interactions have been found for young drivers with gender, speed at yellow and TSL, middle-aged drivers with gender, phone condition, TSL and distance from the stop line at yellow, and older drivers with phone condition and TSL.

The decision tree classified drivers' stop/go decisions through segmenting the dataset into 15 smaller and more homogenous groups. The group statistics that indicate the classification rules for the stop/go decision are presented in the parentheses at the bottom of each branch of the tree in figure 1. The numbers in parenthesis below the leaves indicate how many instances reached the leaf (first number), and of those, how many were not classified as part of the branch (second number). For example, the statistic of branch 1 suggests that 119 novice drivers reached the leaf, and of these, 71 (71/119=59.7%) would stop at the onset of yellow light at signalized intersections. Similarly, branch 2 implies that about 76.3% (29/38) of young females would stop and 23.7% (9/38) of them would proceed through an intersection when the speed at the onset of yellow is less than or equal to 45.5 mph. Branch 12 suggests that about 65% of distracted older drivers would proceed through the intersection when the time to stop line at the onset of yellow is less than or equal to 3.7 seconds.

Using the decision tree, 15 interaction variables were created and shown in third brackets at the bottom of each tree branch. For example, interaction variable 2 refers to a young female driver whose speed at the onset of yellow is less than or equal to 45.5 mph. Similarly, interaction variable 8 refers to a situation where a middle-aged female driver driving in the handheld phone condition and the vehicle's distance from the stop line at the onset of yellow is greater than 162.8 ft and the TSL is less than or equal to 3.72 seconds. Interaction variable 12 refers to a distracted older driver with the TSL at the onset of yellow is less than or equal to 3.7 seconds. These higher order interaction variables coupled with main effect variables and all possible second-order interaction variables were fitted into the repeated measures logistic regression model described previously. Significant variables were retained to derive the parsimonious model.

The significant variables estimated from the logistic regression model along with the probabilities of drivers' decision for proceeding through the intersection at the onset of a yellow light are reported in table 2. The best-fitted model yielded a Wald chi-square statistic of 31.1 with 7 df, which is well above the critical value at 5% significance level, implying that the model has sufficient explanatory power. Moreover, a value of 0.39 for the exchangeable correlation parameter, ρ indicates that there is a significant correlation among observations of each driver and further ensures the appropriateness of the repeated measures logistic regression model for this dataset.

The parsimonious model included three main effect variables: *speed at the onset of yellow light*, *phone condition*, and *driver's gender*, and four higher order interaction variables: *interaction variable 2*, *8*, *10*, and *13* from the decision tree. The parameter

estimates are explained in this section in terms of the odds ratio. That is, each variable's influence on the odds of running a yellow signal is explained while controlling for other effects in the model. However, it is important to recognize that the likelihood of proceeding through a yellow light is a function of many factors, which are described in more detail in the Discussion.

Speed at the onset of yellow was a significant predictor at 5% significance level in explaining driver's decision at the onset of a yellow light at intersections. Results indicate that drivers are more likely to proceed through the intersection with an increased speed. An estimate of odds ratio $[Exp(\beta)]$ indicates that a 1 mph increase in speed is associated with about 4% increase in driver's probability to proceed at the onset of yellow.

The *phone condition* variable suggests that overall, the odds of proceeding through a yellow light decreased by 33% when participants were talking with a handheld phone. It should be noted that the odds ratio estimate does not take into account the effects of interaction variables related to phone condition such as *interaction variable 8* and *interaction variable 13*. Since this study has included a full version of interaction variables, the varied effects of cell phone distraction would be identifiable while interpreting interactions of cell phone distraction with other variables such as driver's age, gender and other situational parameters.

Driver's gender was significant at the 5% level and indicated that female drivers are more likely to run through the yellow light compared to males. The corresponding odds ratio suggests that female drivers are 2.62 times more likely to proceed through a yellow light compared to males.

The *driver's age* factor was grouped into four classes: novice (16-17 years), young (18-25 years), middle-aged (30-45 years) and older (50-60 years). Middle-aged group was used as a reference category for the modeling purpose. None of the main effect variables of age group was significant in the repeated measures logistic regression model. However, the effects of age were evident in the model through four higher order interaction variables derived from the decision tree. While young and older drivers have a higher order interaction each, middle-aged drivers have two higher order interactions in the model.

A higher order interaction variable of young drivers shows that young female driver group has a significant association in driver's stop/go decision model with respect to speed thresholds. *Interaction variable 2* indicates that a young female driver, when drives slower than 45.5 mph, is less likely to proceed through an intersection with the corresponding odds about 70% lower.

Middle-aged female drivers have significant associations in the logistic regression model with respect to time to stop line (TSL), distance from the stop line at the yellow and phone condition. *Interaction variable 8* suggests that a middle-aged female driver driving in cell phone distracted condition is about 71% less likely to proceed through a yellow light when the vehicle's distance at the onset of yellow light is greater than 162.8 ft and the TSL is less than or equal to 3.72 seconds. In contrast, non-distracted driving of a middle-aged female driver appears to have different TSL thresholds in the driver's stop/go logistic regression model. *Interaction variable 10* indicates that a middle-age female driver driving without a cell phone distraction is about 83% less likely to run the yellow light if their distance at the onset of yellow is greater than 162.8 ft and the TSL is in between 2.95 and 3.72 seconds.

Older driver group has a significant association with TSL thresholds in distracted driving condition. *Interaction variable 13* suggests that distracted older drivers are about 6.1 times more likely to run the yellow light if their TSL at the onset of yellow is greater than 3.7 seconds.

To examine the differential risk and behavior arising from distraction and approach speed, a complex model accounting for the interaction terms can be used. To illustrate, we

might be interested in comparing the probabilities of running a yellow signal of distracted older male drivers at approach speeds of 30 and 40 mph. We can calculate the relevant probabilities for each case using logistic regression equation 1. Probabilities of yellow light running of a discrete group of individuals at different approach speed values could be computed by using corresponding $\mathbf{X}$ values and parameter estimates, while holding other variables at their reference category values. Relevant parameter estimates for distracted older male drivers are *constant*, *speed at yellow*, *handheld phone* and *Interaction variable 13*. Taking the corresponding parameter estimates from table 2, the predicted probabilities for this group of driver at approach speed = 30 mph and approach speed = 40 mph are calculated as follows:

$$P_{\text{Distracted Older Male Speed}=30} = \frac{\text{EXP}[-2.12 + 0.04(30) - 0.406(1) + 1.813]}{1 + \text{EXP}[-2.12 + 0.04(30) - 0.406(1) + 1.813]} = 0.62$$

$$P_{\text{Distracted Older Male Speed}=40} = \frac{\text{EXP}[-2.12 + 0.04(40) - 0.406(1) + 1.813]}{1 + \text{EXP}[-2.12 + 0.04(40) - 0.406(1) + 1.813]} = 0.71$$

Using this approach for a variety of driver groups, figure 2 presents the model predicted probabilities of running a yellow light as a function of approach speed. Since observed approach speeds of drivers ranged between 25 and 55 mph, the yellow light running probabilities are plotted for that speed range. Findings from this plot have been elaborately discussed in the next section.

5. Discussion

On average and without the effects of interaction variables, participants in this study are more likely to stop at the onset of a yellow light when engaged in a cell phone conversation (odds ratio 0.666). This effect might be explained by risk compensation—where the sampled drivers' compensated for the increased risk induced by distraction by reducing their willingness to run the yellow light. During a conversation, a driver may have reduced mental capacity to attend to the driving task, which could be perceived as a reduction in the ability to respond quickly to a hazardous situation. It is also possible that the increased workload of talking on the phone results in resources allocated away from other tasks, one of them being the assessment of the risks of running the yellow light. While this effect may appear to be risk compensation, it could simply be the result of cognitive resource reallocation. From this study, it is not possible to know whether risk compensation or cognitive resource allocation has resulted in reduced yellow light running. There are a number of studies that associate crash risks with engagement in cell phones while driving (e.g., Neyens & Boyle, 2007; Redelmeier & Tibshirani, 1997). Hence, if risk compensation is occurring, it is likely to be insufficient to offset increased crash risk caused by distraction.

Risk compensation of distracted drivers has been repeatedly noted in the literature as a possible explanation for observed speed reductions (e.g., Haigney, Taylor, & Westerman, 2000; Törnros & Bolling, 2006). It has generally been argued that drivers tend to compensate for the increase workload of talking on their cell phone whilst driving by selecting a lower driving speed. To test whether any speed reduction was evident in this dataset, drivers' approach speeds to a signalized intersection were tested and compared across phone conditions using the repeated measures ANOVA in the form of a Linear Mixed Model as used by Haque and Washington (2013). As shown in table 3, the difference in driving speed at the onset of yellow—measured at the instance when the traffic light changed from green to yellow—was statistically significant ($F_{1, 339.84} = 5.493$, $p\text{-value} = 0.020$) across phone conditions. The mean speed at yellow in baseline driving was 43.2 mph, while the

1 corresponding speed in cell phone conversation condition was 42.2. On average, drivers
2 tended to drive 1 mph slower while driving distracted.

3 Overall, the results suggest that drivers tend to compensate for their increased
4 attention load of cell phone conversation by not only reducing the probability of yellow light
5 running but also selecting a lower speed while approaching to a signalized intersection. In
6 contrast, other research has shown that drivers engaged in a phone conversation are more
7 likely to miss (Strayer & Johnston, 2001) or react slower to critical signals (Consiglio et al.,
8 2003) and changing stop lights (Hancock et al., 2003). Drivers' responses to cell phone
9 distraction and changes in traffic lights appear to behave in complex ways, as is evident from
10 many higher order interactions in the logistic regression model. A closer examination of
11 driver behavior mainly comparing yellow running probabilities across different driver groups
12 (e.g. young male, older female, etc.) is provided later in this section.

13 The model suggests that females are less likely to stop for a yellow signal. The
14 average approach speed of female drivers at the onset of yellow light was 42.2 mph, while the
15 corresponding speed for male drivers was 42.7 mph. The speed difference between males and
16 females is not statistically significant ($p\text{-value} = 0.313$). This insignificant difference in speed
17 combined with females being less likely to stop implies that females may be less responsive
18 to traffic light changes when compared to males traveling at similar speeds.

19 The risk of running a yellow signal increases by a factor of 4% (odds ratio 1.04) for
20 those who travel faster than the mean speed by 1 mph. Similar findings were also observed in
21 other studies (e.g., Papaioannou, 2007) where drivers were more likely to proceed through an
22 intersection when their approach speeds were higher.

23 The presence of high order interaction effects indicates some complexity in the
24 relationships between driver's stop/go decision at the onset of yellow and operating speed,
25 distance from the stop line, time to stop line (TSL), phone condition and driver demographics
26 including age and gender. The effects of cell phone distraction on the speed selection of
27 drivers approaching signalized intersections have further complicated these relationships. To
28 examine the differential risk, probabilities of yellow light running across various driver
29 groups are plotted in figure 2.

30 As shown in figure 2, the probability of yellow light running across all age groups
31 generally tends to increase with the approach speed at the onset of yellow light. For example,
32 the probability of yellow light running for non-distracted male drivers at 25 mph is 25%,
33 while at an approach speed of 55 mph the corresponding probability is 42%. Both males and
34 females have a lower probability of yellow light running across the whole speed range when
35 they are engaged in a cell phone conversation while driving. For instance, the probability of
36 yellow running of female drivers at 45 mph is about 10% lower in distracted condition
37 compared to non-distracted driving.

38 Distracted female drivers, however, have a higher probability of yellow light running
39 across the whole speed range compared to distracted male drivers. The probability of yellow
40 running at 45 mph, for instance, for distracted female drivers is 56% while the corresponding
41 probability for distracted males is only 33%. Speed selection of female drivers also seems to
42 be influenced by the cell phone conversation task. As reported in table 3, the difference in
43 driving speed across phone conditions was not statistically significant ($F_{1, 227.76} = 1.365$, $p\text{-value} = 0.245$) for male drivers but it was significant ($F_{1, 110.95} = 7.366$, $p\text{-value} = 0.008$) for
44 female drivers. While the driving speed of non-distracted female drivers at the onset of
45 yellow was about 43.4 mph, the corresponding speed for distracted females was about 41.6
46 mph. In summary, the effects of distraction in terms of both yellow light running and speed
47 reduction appear to be higher among female drivers than males.

48 Both young male and female drivers appear to reduce their yellow light running
49 probabilities in the distracted driving condition across the whole speed range, with the
50

1 corresponding probability on average about 10% lower in the distracted condition compared
2 to normal driving. Interestingly the probability of yellow light running of young females in
3 both distracted and non-distracted conditions increases significantly, when the approach
4 speed exceeds the posted speed limit of 45 mph. At an approach speed of 45 mph, the
5 probability of yellow light running of distracted female drivers is 28% while the
6 corresponding probability at the approach speed of 50 mph is 61%. Drivers travelling faster
7 than the speed limits or the speed values used for determining the traffic signal time settings
8 are likely to be caught in a dilemma zone, where they can neither cross the intersection
9 without red light running nor stop without hard braking (Papaioannou, 2007). The effect of
10 the dilemma zone appears to have a greater influence on young female drivers as their
11 probability of yellow running is significantly increased when their approach speed exceeds
12 the speed limit.

13 Like young males, middle-age male drivers have a similar trend of yellow light
14 running, with a similar reduction in yellow running probabilities in distracted condition.
15 However, the difference of yellow running probabilities in distracted and baseline driving of
16 middle-aged female drivers is marginal across the whole speed range. On average, middle-
17 aged female drivers have a lower probability of yellow running compared to middle-aged
18 males.

19 In contrast, older drivers appear to have an opposite pattern in yellow light running
20 risk. Both older males and females have a higher probability of yellow light running in
21 distracted driving condition compared to the baseline across the whole speed range. For
22 instance, the probability of yellow light running at the approach speed of 45 mph for a non-
23 distracted older male driver is about 42% while the corresponding probability for a distracted
24 older male is as high as 75%. On average, the probability of yellow light running of older
25 males is about 33% higher in the distracted driving condition and the difference is constant
26 across the whole speed range. Older female drivers exhibit the same effect, and it marginally
27 decrease with increasing speed; that is the probability difference decreases with approach
28 speed. At low approach speeds (say 30 mph), the probability of yellow running of older
29 females is about 30% higher in distracted condition compared to baseline driving, while the
30 corresponding probability difference at an approach speed 55 mph is about 18%. Previous
31 research (e.g., Knoblauch et al., 1995) has documented that average slower response times of
32 older drivers might impair their actions at the onset of yellow light. This effect might be
33 exacerbated by a cell phone distraction as evident in the yellow light probabilities of
34 distracted older drivers. Apart from this, speed selection of older drivers along an approach of
35 a signalized intersection also tends to be significantly ($F_{1, 66.00} = 6.131$, $p\text{-value} = 0.016$)
36 affected by the cell phone conversation task (see table 3). Approach speed of older driver in
37 baseline driving condition was about 43.1 mph while the corresponding speed in distracted
38 condition was about 40.6 mph.

39 Compared with older male drivers, probabilities of yellow running across the entire
40 speed range are higher among older females both in distracted and non-distracted driving
41 conditions. For instance, the probability of yellow running at 45 mph for distracted older
42 males is 75%, while the corresponding probability for older females is as high as 89%. In
43 addition, the cell phone distraction appears to have a significant effect on the speed selection
44 of older females ($F_{1, 25.69} = 18.312$, $p\text{-value} < 0.001$) but not among older males ($F_{1, 39.00} =$
45 0.833 , $p\text{-value} = 0.367$). Average approach speeds along an approach of a signalized
46 intersection for older females in distracted and baseline driving were 44.9 and 40.5 mph,
47 respectively. That is the speed reduction was about 4.4 mph due to the cell phone
48 conversation task. In summary, the influence of a cell phone conversation on driving
49 performance of older female driver group along an approach to signalized intersection
50 appears to be the most among all driver groups.

6. Conclusions

This study applied a combination of data mining and classical statistical modeling techniques to examine distracted drivers' stop/go decisions at the onset of yellow lights at signalized intersections. Data were obtained from a group of drivers on a motion-based driving simulator (NADS-1). The methodology combined logistic regression and decision trees in order to leverage the strengths of both approaches—an ability to test significance of observed effects with an ability to heuristically explore the data for previously unknown relationships. The combined methodology provided useful insights regarding interaction effects that significantly influenced drivers' decisions to proceed through a yellow light while distracted by a cell phone conversation.

Overall, the model suggests that driver responses to distraction and running a yellow signal vary considerably across gender and age. Female drivers have a higher probability of running through a yellow light in general. Risk compensation, or complexity compensation, was observed in young and middle-aged driver groups. Distracted young and middle-aged drivers are predicted to compensate for the increased mental load and/or crash risk of distraction by running the yellow signal with lower probability. It is not clear from this study whether drivers were responding to increased perception of risk or simply an increased cognitive workload. However, yellow light running is associated with crash risk, and this study showed that distracted young and middle-aged drivers reveal a reduction in their willingness to accept as much risk as when not distracted. In contrast, older drivers appear to behave quite differently in the presence of distraction in this scenario and predicted by this model. Overall, older drivers have a higher probability of yellow light running while distracted by a cell phone conversation. In particular, the performance of older females appears to be the most affected by the cell phone distraction.

Speed selection of drivers on an approach to a signalized intersection also appears to be influenced by the cell phone conversation task. In general, drivers tended to select lower speeds when they were engaged in cell phone conversations. In particular, female drivers tended to reduce their driving speeds while distracted by the cell phone conversations, unlike males. Distracted older drivers tended to select slower approach speeds to traffic signals. Hence, the effects of distraction on speed reduction as well as yellow running appear to be higher for these two groups of drivers.

This study was based on a reanalysis of existing data and as such, there were some limitations given that the original study goals were not aligned with the aims of this paper. The analysis was based on two different studies (albeit the same protocol), there may be differences between novice drivers and the other three age groups that may not be related to age, but rather some subtle differences in each study goal (one was specifically on use of wireless devices, while the other was on novice driver risk propensity). Furthermore, the novice drivers consisted of only males and as such, any differences in gender for novice drivers could not be detected. It is also recognized that drivers may exhibit different responses between initial and final exposure to a yellow light given learning effects, which may also bias outcomes. Hence, additional simulator studies would be useful to capture the changes in perception and workload, as well as to identify and test ways to mitigate the effects of driver distraction at intersections.

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TABLE 1 Summary statistics of variables included in the model

Variables	Count on encounters with traffic light change	Mean	St. Dev.
<i>Driver Demographics</i>			
Age			
Novice	119	-	-
Young	108	-	-
Middle*	102	-	-
Older	80	-	-
Gender			
Male*	275	-	-
Female	134	-	-
<i>Experiment Variables</i>			
Phone Condition			
Baseline*	136	-	-
Handheld	273	-	-
Speed at yellow (mph)	-	42.52	5.17
Time-to-stop line (seconds)	-	3.27	0.37
Distance at yellow (ft)	-	204.22	33.76
Drivers' decision to proceed through a yellow light	153	-	-

*Reference category for categorical independent variables

TABLE 2 Repeated measure logistic regression model predicting the likelihood of proceeding through a yellow light

Explanatory Variables	Parameter Estimate		z-statistic	p-value	Odds Ratio (OR)	95% CI of OR	
	Estimate	SE				2.5%	97.5%
Speed at yellow	0.040	0.016	2.54	0.011	1.041	1.009	1.073
Handheld Phone	-0.406	0.143	-2.84	0.004	0.666	0.504	0.882
Female	0.963	0.420	2.29	0.022	2.620	1.150	5.968
Interaction variable 2	-1.185	0.483	-2.45	0.014	0.306	0.119	0.788
Interaction variable 8	-1.264	0.656	-1.93	0.054	0.283	0.078	1.023
Interaction variable 10	-1.740	0.740	-2.35	0.019	0.175	0.041	0.748
Interaction variable 13	1.813	0.847	2.14	0.032	6.131	1.167	32.219
Constant	-2.120	0.758	-2.80	0.005			
Number of observations	409						
Number of groups	69						
Wald Chi-sq	31.08						
Degrees of freedom	7						
p-value	< 0.001						
Exchangeable correlation parameter, ρ	0.3860						

Interaction variable 2: A young female driver whose speed at the onset of yellow is ≤ 45.5 mph
Interaction variable 8: A distracted middle aged female driver whose distance from the stop line at yellow is > 162.8 ft and time to stop line is ≤ 3.72 seconds
Interaction variable 10: A middle aged female driver without phone conversations whose distance from the stop line at yellow is > 162.8 ft and time to stop line is between 2.95 and 3.72 seconds
Interaction variable 13: A distracted older driver whose time to stop line at the onset of yellow is > 3.7 seconds

TABLE 3 Speed selection of different driver groups at the onset of yellow light

Speed at Yellow	Phone Condition		Significance by a Linear Mixed Model	Remark
	Baseline	Handheld		
All Drivers	43.212	42.170	$F_{1, 339.84} = 5.493, p = 0.020$	Significant
Age				
Novice	42.914	42.261	$F_{1, 97.68} = 0.558, p = 0.457$	Not significant
Young	44.017	43.285	$F_{1, 89.00} = 0.948, p = 0.333$	Not significant
Middle	42.762	42.163	$F_{1, 84.00} = 0.485, p = 0.488$	Not significant
Older	43.142	40.559	$F_{1, 66.00} = 6.131, p = 0.016$	Significant
Gender				
Male	43.123	42.471	$F_{1, 227.76} = 1.365, p = 0.245$	Not significant
Female	43.398	41.558	$F_{1, 110.95} = 7.366, p = 0.008$	Significant
Age X Gender				
Novice Male	42.914	42.261	$F_{1, 97.68} = 0.558, p = 0.457$	Not significant
Young Male	45.013	43.820	$F_{1, 44.00} = 1.003, p = 0.322$	Not significant
Young Female	43.021	42.750	$F_{1, 44.00} = 0.086, p = 0.771$	Not significant
Middle Male	42.640	43.218	$F_{1, 44.00} = 0.289, p = 0.594$	Not significant
Middle Female	42.898	40.975	$F_{1, 39.00} = 2.024, p = 0.163$	Not significant
Older Male	42.060	40.633	$F_{1, 39.00} = 0.833, p = 0.367$	Not significant
Older Female	44.874	40.453	$F_{1, 25.69} = 18.312, p < 0.001$	Significant

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FIGURE 1 Decision tree diagram for the stop/go decision model

FIGURE 2 Probability of proceeding through yellow light for different driver groups